

# The distribution of the maximum of character sums

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## Background & motivation

Let  $\chi$  be a Dirichlet character modulo  $q$  and define

$$M(\chi) = \max_{1 \leq z \leq q} \left| \sum_{n \leq z} \chi(n) \right|.$$

If  $\chi$  is non-principal, then [Pólya and Vinogradov](#) showed in 1918 that

$$M(\chi) \ll \sqrt{q} \log q.$$

Assuming GRH, [Montgomery and Vaughan](#) improved this in 1977 to

$$M(\chi) \ll \sqrt{q} \log \log q.$$

This is best possible: [Paley](#) had already shown in 1932 that

there is a sequence  $q_n \rightarrow \infty$  such that  $M\left(\left(\frac{\cdot}{q_n}\right)\right) \gg \sqrt{q_n} \log \log q_n$ .

However, such extremal examples should be rather rare. Our goal is to study *how rare* they are.

## The distribution of $M(\chi)$ : random models

We shall study

$$P_q(\tau) := \frac{\#\{\chi \pmod{q} : M(\chi) > \frac{e^\gamma}{\pi} \tau \sqrt{q}\}}{\phi(q)} = \mathbf{Prob} \left( M(\chi) > \frac{e^\gamma}{\pi} \tau \sqrt{q} \right).$$

We always assume for simplicity that  $q$  is prime.

Two questions:

- 1 Is there a random model that describes  $P_q(\tau)$  accurately?
- 2 How big is  $P_q(\tau)$ ?

Let  $(X_p)_{p \nmid q}$  be a sequence of independent random variables, uniformly distributed on  $\{z \in \mathbb{C} : |z| = 1\}$ , and  $X_p = 0$  if  $p|q$ .

(They should model  $\chi(p)$  as  $\chi$  runs through characters modulo  $q$ .)

Then we define  $X_n = \prod_{p^r \parallel n} X_p^r$ , which serves a model for  $\chi(n)$ .

**First attempt:** model  $\sum_{n \leq z} \chi(n)$  by  $\sum_{n \leq z} X_n$ .

For  $z$  large compared to  $q$ , this will fail: periodicity is not taken into account.

## The distribution of $M(\chi)$ : random models, continued

$$P_q(\tau) = \mathbf{Prob} \left( M(\chi) > \frac{e^\gamma}{\pi} \tau \sqrt{q} \right).$$

$(X_p)_{p|q}$  sequence of independent random variables, uniformly distributed on  $\{z \in \mathbb{C} : |z| = 1\}$ ,  $X_p = 0$  if  $p \nmid q$ ,  $X_n = \prod_{p^r || n} X_p^r$ .

**Second attempt:** use Pólya's expansion ( $\chi$  primitive,  $e(x) = e^{2\pi i x}$ ):

$$\sum_{n \leq z} \chi(n) = \frac{\tau(\chi)}{2\pi i} \sum_{1 \leq |n| \leq w} \frac{\bar{\chi}(n)(1 - e(-nz/q))}{n} + O\left(\frac{q \log q}{w}\right) \quad (1 \leq w \leq q).$$

Our model for  $\sum_{n \leq z} \chi(n)$  then becomes

$$S(z) := \frac{\tau(\chi)}{2\pi i} \sum_{1 \leq |n| \leq q} \frac{\bar{X}_n \cdot (1 - e(-nz/q))}{n}.$$

This model captures the periodicity of  $\chi$ .

**Remark.** The standard deviation of  $S(z)/\sqrt{q}$  is  $\ll 1$ . (Compare this to 1st model: the SD of  $T(z) = \sum_{n \leq z} X_n$  is  $\sqrt{z}$ , and one might expect  $T(z)/\sqrt{z}$  to get large relatively often.) As a result,  $P_q(\tau)$  will be rather small.

## Known results on $P_q(\tau)$

In 1979, [Montgomery and Vaughan](#) showed that

$$\frac{1}{\phi(q)} \sum_{\chi \pmod{q}} M(\chi)^{2k} \ll_k q^k.$$

An immediate corollary is that

$$P_q(\tau) = \mathbf{Prob} \left( M(\chi) > \frac{e^\gamma}{\pi} \tau \sqrt{q} \right) \ll_A \frac{1}{\tau^A}.$$

In 2011, [Bober-Goldmakher](#) proved that, for fixed  $\tau$  and  $q \rightarrow \infty$  over primes,

$$\exp \left\{ -\frac{C e^\tau}{\tau} (1 + o_{\tau \rightarrow \infty}(1)) \right\} \leq P_q(\tau) \leq \exp \left\{ -e^{B\sqrt{\tau}/(\log \tau)^{1/4}} \right\},$$

where  $C = 1.09258\dots$ . This supports the claim that  $P_q(\tau)$  is very small.

**Question:** why do the tails of the distribution of  $M(\chi)$  have this double exponential decay?

## The distribution of $M(\chi)$ vs the distribution of $L(1, \chi)$

$$\sum_{n \leq \alpha q} \chi(n) = \frac{\tau(\chi)}{2\pi i} \sum_{1 \leq |n| \leq q} \frac{\bar{\chi}(n)(1 - e(-n\alpha))}{n} + O(\log q).$$

In view of Pólya's expansion, one might conjecture that

$$P_q(\tau) := \mathbf{Prob} \left( M(\chi) > \frac{e^\gamma}{\pi} \tau \sqrt{q} \right) \approx \mathbf{Prob} (|L(1, \chi)| > c\tau).$$

**Granville-Soundararjan:** for  $q$  prime and  $e^\tau = o(\log q)$ ,

$$\mathbf{Prob} (|L(1, \chi)| > e^\gamma \tau) = \exp \left\{ -\frac{Ce^\tau}{\tau} (1 + o_{\tau \rightarrow \infty}(1)) \right\}.$$

Compare this to

$$\exp \left\{ -\frac{Ce^\tau}{\tau} (1 + o_{\tau \rightarrow \infty}(1)) \right\} \leq P_q(\tau) \leq \exp \left\{ -e^{B\sqrt{\tau}/(\log \tau)^{1/4}} \right\}.$$

## The distribution of $L(1, \chi)$ : main ideas

- ① We shall take moments of  $L(1, \chi)$ , so we need to 'shorten' it. We have that  $\log L(1, \chi) = \sum_p \chi(p)/p + C_\chi$ , where  $C_\chi$  is a constant.

$$\text{PNT} \Rightarrow \log L(1, \chi) \sim \sum_{p \leq e^{q^\epsilon}} \chi(p)/p + C_\chi,$$

$$\text{GRH} \Rightarrow \log L(1, \chi) \sim \sum_{p \leq (\log q)^{2+\epsilon}} \chi(p)/p + C_\chi.$$

But we study  $L(1, \chi)$  *statistically*: for most  $\chi \pmod{q}$ ,

Zero-density estimates  $\Rightarrow \log L(1, \chi) \sim \sum_{p \leq (\log q)^{100}} \chi(p)/p + C_\chi$ .

- ② Take moments of

$$L(1, \chi; y) := \prod_{p \leq y} \left(1 - \frac{\chi(p)}{p}\right)^{-1} = \sum_{p|n \Rightarrow p \leq y} \frac{\chi(n)}{n} \quad (y = (\log q)^{100}) :$$

$$\begin{aligned} \frac{1}{\phi(q)} \sum_{\chi \pmod{q}} |L(1, \chi; y)|^{2k} &= \frac{1}{\phi(q)} \sum_{\chi \pmod{q}} \left| \sum_{p|n \Rightarrow p \leq y} \frac{\tau_k(n) \chi(n)}{n} \right|^2 \\ &= \sum_{\substack{m \equiv n \pmod{q} \\ p|mn \Rightarrow p \leq y, p \nmid q}} \frac{\tau_k(m) \tau_k(n)}{mn}. \end{aligned}$$

## The distribution of $L(1, \chi)$ , continued

Ignoring the off-diagonal terms (assumption that  $X_n$  is a good model for  $\chi(n)$ ), and assuming that  $q$  is prime,

$$M_{2k} := \frac{1}{\phi(q)} \sum_{\chi \pmod{q}} |L(1, \chi; y)|^{2k} = \sum_{\substack{m \equiv n \pmod{q} \\ p|mn \Rightarrow p \leq y, p \nmid q}} \frac{\tau_k(m)\tau_k(n)}{mn}$$
$$\approx \sum_{p|n \Rightarrow p \leq y} \frac{\tau_k(n)^2}{n^2} = \prod_{p \leq y} \left( 1 + \frac{\tau_k(p)^2}{p^2} + \frac{\tau_k(p^2)}{p^4} + \dots \right).$$

Then Granville and Soundararajan proceed to show that  $\log M_{2k} = 2e^\gamma k + C'k/\log k + O(k/\log^2 k)$ , which allows them to estimate  $\mathbf{Prob}(|L(1, \chi)| > e^\gamma \tau)$  quite accurately.

**Remark.** In fact, they observe that

$$\log \left( 1 + \frac{\tau_k(p)^2}{p^2} + \frac{\tau_k(p^2)}{p^4} + \dots \right) = \log I_0(2k/p) + O(k/p^2),$$

where  $I_0(t) = \sum_{n \geq 0} \left( \frac{t/2}{n!} \right)^2$  is the modified Bessel function of the 1st kind. In particular, most of the contribution to  $M_{2k}$  comes from primes  $p \approx k$ .



## New results on $P_q(\tau) = \mathbf{Prob} \left( M(\chi) > \frac{e^\gamma}{\pi} \tau \sqrt{q} \right)$

Recall Bober-Goldmakher's result: for  $\tau$  fixed and  $q \rightarrow \infty$  over primes,

$$\exp \left\{ -\frac{Ce^\tau}{\tau} (1 + o_{\tau \rightarrow \infty}(1)) \right\} \leq P_q(\tau) \leq \exp \left\{ -e^{B\sqrt{\tau}/(\log \tau)^{1/4}} \right\}.$$

There are two issues to be addressed:

- There is a discrepancy between upper and lower bounds.
- The result is not uniform in  $\tau$  and  $q$ .

### Theorem (Bober, Goldmakher, Granville, K. (2013))

Let  $\theta > 14/15$ ,  $q$  be prime and  $2 \leq \tau \leq \log \log q - \log \log \log q - 5$ . Then

$$\exp \left\{ -\frac{Ce^\tau}{\tau} (1 + o_{\tau \rightarrow \infty}(1)) \right\} \leq P_q(\tau) \leq \exp \left\{ -e^{\tau + O_\theta(\tau^\theta)} \right\}.$$

**Remark.** On GRH, the theorem holds when  $\tau \leq \log_2 q - \log_4 q + O(1)$ . It seems likely that it can be shown unconditionally  $e^\tau = o(\log q)$  can be obtained unconditionally.

## A reduction to the distribution of $L(1, \chi)$ : lower bounds

For lower bounds on  $P_q(\tau)$ , we follow Bober-Goldmakher and note that

$$\begin{aligned} \sum_{n \leq q/2} \chi(n) &\sim \frac{\tau(\chi)}{2\pi i} \sum_{1 \leq |n| \leq q} \frac{\bar{\chi}(n)(1 - e(-n/2))}{n} \\ &= \frac{\tau(\chi)}{\pi i} \sum_{\substack{1 \leq |n| \leq q \\ n \text{ odd}}} \frac{\bar{\chi}(n)}{n} = \frac{2\tau(\chi)}{\pi i} \begin{cases} \sum_{\substack{1 \leq n \leq q \\ n \text{ odd}}} \frac{\bar{\chi}(n)}{n} & \text{if } \chi(-1) = -1, \\ 0 & \text{if } \chi(-1) = 1. \end{cases} \end{aligned}$$

When  $\chi$  is odd, the right hand side is essentially  $L(1, \bar{\chi})$ , divided by the Euler factor at  $p = 2$ .

One can then obtain the claimed lower bound on  $P_q(\tau)$  using the methods of Granville-Soundararajan.

The upper bound is significantly harder. The main issue is to understand where  $\sum_{n \leq z} \chi(n)$  is maximized (ideas about pretentious characters).

## A detour: pretentious characters

Granville-Soundararajan (2006) and Goldmakher (2010) improved the previously known bounds for  $M(\chi)$  when  $\chi$  has odd order  $g$  to

$$M(\chi) \ll \begin{cases} \sqrt{q}(\log q)^{1-\delta_g+o(1)} & \text{unconditionally,} \\ \sqrt{q}(\log \log q)^{1-\delta_g+o(1)} & \text{on GRH.} \end{cases} \quad \left( \delta_g = 1 - \frac{g}{\pi} \sin \frac{\pi}{g} \right)$$

**Idea of the proof:**  $g$  odd  $\Rightarrow \chi(-1) = 1$ . So Pólya's expansion becomes

$$\sum_{n \leq \alpha q} \chi(n) \sim \frac{\tau(\chi)}{2\pi i} \sum_{1 \leq |n| \leq q} \frac{\bar{\chi}(n)(1 - e(-n\alpha))}{n} = -\frac{\tau(\chi)}{2\pi i} \sum_{1 \leq |n| \leq q} \frac{\bar{\chi}(n)e(-n\alpha)}{n}.$$

Let  $|\alpha - a/b| < 1/(bB)$ ,  $b \leq B := e^{\sqrt{\log q}}$ . Montgomery-Vaughan showed

$$\sum_{n \leq x} \frac{\chi(n)e(n\alpha)}{n} \ll \log \log x + \log b + \frac{(\log b)^{3/2}}{\sqrt{b}} \log x \quad (x \geq 2).$$

So, we may assume that  $b \leq (\log q)^{1/3}$ . Also, let  $\alpha = a/b$  for simplicity.

## Pretentious characters, continued

$\text{ord}(\chi) = g = \text{odd}$ ,  $\chi(-1) = 1$ ,  $\alpha = a/b$ ,  $b \leq (\log q)^{1/3}$ . We need to estimate

$$\sum_{n \leq \alpha q} \chi(n) \sim -\frac{\tau(\chi)}{2\pi i} \sum_{1 \leq |n| \leq q} \frac{\bar{\chi}(n)e(-na/b)}{n}.$$

Expand  $e(-na/b)$  in terms of characters  $\psi \pmod{d}$ ,  $d|b$ , to replace  $\sum_{n \leq \alpha q} \chi(n)$  by sums of the form

$$S = \sum_{1 \leq |n| \leq z} \frac{\bar{\chi}(n)\psi(n)}{n} = (1 - \chi(-1)\psi(-1)) \sum_{n \leq z} \frac{\bar{\chi}(n)\psi(n)}{n}.$$

For  $S$  to be big,  $\chi$  must be 'close' to  $\psi$  ( $\chi(p) \approx \psi(p)$ ). Indeed,

$$\sum_{n \leq z} \frac{\chi(n)\bar{\psi}(n)}{n} \ll \frac{\log z}{\exp\{\mathbb{D}(\chi, \psi; z)/2\}}, \quad \mathbb{D}^2(\chi, \psi; z) = \sum_{p \leq z} \frac{1 - \Re(\chi(p)\bar{\psi}(p))}{p}.$$

If  $\mathbb{D}(\chi, \psi; z)$  is small, we say that  $\chi$  *pretends to be*  $\psi$ .

Also,  $\psi(-1) = -\chi(-1) = -1 \implies \text{ord}(\psi) = \text{even} \neq g$ .

But then  $\chi \approx \psi \implies 1 = \chi^g \approx \psi^g \neq 1$ , a contradiction.

## A reduction to the distribution of $L(1, \chi)$ : upper bounds

In bounding  $P_q(\tau)$  from above, the key step is the following:

**Theorem (Bober, Goldmakher, Granville, K. (2013))**

Let  $\theta > 14/15$ ,  $q$  be prime and  $2 \leq \tau \leq \log \log q - \log \log \log q - 5$ . With the exception of  $\ll q \exp\{-20\tau e^\tau\}$  characters mod  $q$ , if  $M(\chi) > \frac{e^\gamma}{\pi} \tau \sqrt{q}$ , then  $\chi$  is odd, and there is a  $b \leq \tau^{10}$  such that

$$\left| \sum_{\substack{n \in \mathbb{N}, (n, b) = 1 \\ p|n \Rightarrow p \leq e^\tau}} \frac{\chi(n)}{n} \right| \geq e^\gamma \tau + O_\theta(\tau^\theta).$$

Then  $P_q(\tau) \leq \exp\{-e^{\tau + O_\theta(\tau^\theta)}\}$ , by Granville-Soundararajan.

Main ideas involved in proving the above theorem:

- 1 A high moment bound to truncate Pólya's expansion.
- 2 Use "pretentious characters" to locate the max of  $|\sum_{n \leq x} \chi(n)|$ .
- 3 Slow variance of  $\sum_{n \leq x} \chi(n)$  (Lipschitz bounds).

## Truncating Pólya's expansion

When  $\chi$  is primitive, we have that

$$M(\chi) = \max_{\alpha \in [0,1]} \left| \sum_{n \leq \alpha q} \chi(n) \right| = \frac{\sqrt{q}}{2\pi} \max_{\alpha \in [0,1]} \left| \sum_{1 \leq |n| \leq q} \frac{\chi(n)(1 - e(n\alpha))}{n} \right|.$$

Using a moments argument, we show that, for most  $\chi$ ,

$$\sum_{1 \leq |n| \leq q} \frac{\chi(n)(1 - e(n\alpha))}{n} \sim \sum_{1 \leq |n| \leq q, P^+(n) \leq y} \frac{\chi(n)(1 - e(n\alpha))}{n},$$

with  $y \approx e^\tau$  (here  $P^+(n) = \max\{p|n\}$  and  $P^-(n) = \min\{p|n\}$ ). This is done by observing that their difference equals

$$\begin{aligned} \sum_{\substack{1 \leq |n| \leq q \\ P^+(n) > y}} \frac{\chi(n)(1 - e(n\alpha))}{n} &= \sum_{\substack{1 \leq |g| \leq q \\ P^+(g) \leq y}} \frac{\chi(g)}{g} \sum_{\substack{y < h \leq q/g \\ P^-(h) > y}} \frac{\chi(h)(1 - e(gh\alpha))}{h} \\ &\ll \sum_{P^+(g) \leq y} \frac{1}{g} \max_{\alpha \in [0,1]} \left| \sum_{y < h \leq q/g, P^-(h) > y} \frac{\chi(h)e(h\alpha)}{h} \right|. \end{aligned}$$

## Truncating Pólya's expansion, continued

$$\sum_{\substack{1 \leq |n| \leq q \\ P^+(n) > y}} \frac{\chi(n)(1 - e(n\alpha))}{n} \ll \sum_{P^+(g) \leq y} \frac{1}{g} \max_{\alpha \in [0,1]} \left| \sum_{y < h \leq q/g, P^-(h) > y} \frac{\chi(h)e(h\alpha)}{h} \right|.$$

We raise both sides to  $2k$ . Then  $\max_{\alpha \in [0,1]}$  is removed by noticing that  $|\alpha - r/R|$  for some  $r \in \{1, \dots, R\}$ . It remains to estimate

$$\sum_{\chi(\bmod q)} \left| \sum_{y < h \leq q/g, P^-(h) > y} \frac{\chi(h)e(hr/R)}{h} \right|^{2k}.$$

Then we find that this is  $\lesssim \sum_{P^-(n) > y, n > y^k} \tau_k(n)^2/n^2 = o(1)$  if  $k \leq y/(\log y)^{100}$ . (If  $y > k$ , the primes  $p \approx k$  that give most of the contribution to the sum  $\sum_{n \geq 1} \tau_k(n)^2/n^2$  are not present.)

$$y \approx e^\tau \Rightarrow P_q(\tau) \sim \mathbf{Prob} \left( \max_{\alpha \in [0,1]} \left| \sum_{P^+(|n|) \leq e^\tau} \frac{\chi(n)(1 - e(n\alpha))}{n} \right| > 2e^\gamma \tau \right).$$

## Locating the maximum

$$P_q(\tau) \sim \mathbf{Prob} \left( \max_{\alpha \in [0,1]} \left| \sum_{P^+(|n|) \leq e^\tau} \frac{\chi(n)(1 - e(n\alpha))}{n} \right| > 2e^\gamma \tau \right).$$

Write  $N(\chi)$  for the above maximum, and let  $\alpha_\chi$  be its location.

Let  $|\alpha_\chi - a/b| < 1/(bB)$ ,  $b \leq B := e^{\sqrt{\tau}}$ . Also, let  $\xi$  be the primitive character of conductor  $\leq \tau$  that lies the 'closest' to  $\chi$ , i.e.

$$\mathbb{D}(\chi, \xi; e^\tau) = \min_{\substack{\psi \bmod d \leq \tau \\ \psi \text{ prim.}}} \mathbb{D}(\chi, \psi; e^\tau), \quad \mathbb{D}^2(f, g; y) = \sum_{p \leq y} \frac{1 - \Re(f(p)\bar{g}(p))}{p}.$$

**Claim:** If  $N(\chi) > 2e^\gamma \tau$ , then  $\xi = 1$  and  $\chi$  is odd.

Assume not. Then

$$\sum_{P^+(|n|) \leq e^\tau} \frac{\chi(n)}{n} = (1 - \chi(-1)) \sum_{P^+(|n|) \leq e^\tau} \frac{\chi(n)}{n} = o(\tau).$$

Also,  $b \leq \tau^{1/10}$ ; else,  $N(\chi) \sim \sum_{P^+(|n|) \leq e^\tau} \chi(n)e(n\alpha)/n = o(\tau)$ , by [Montgomery-Vaughan](#), a contradiction to " $N(\chi) > 2e^\gamma \tau$ ".



## Locating the maximum, continued

If  $\xi \pmod{D}$  is the 'closest' character to  $\chi$ , and either  $\xi \neq 1$  or  $\chi$  is even:

$$2e^{\gamma\tau} < N(\chi) \sim \left| \sum_{P^+(|n|) \leq e^\tau} \frac{\chi(n)e(n\alpha)}{n} \right|, \quad |\alpha_\chi - a/b| \leq \frac{1}{be\sqrt{\tau}}, \quad b \leq \tau^{1/10}.$$

Assume that  $\alpha_\chi = a/b$ , and expand  $e(na/b)$  using characters, to get sums

$$\sum_{P^+(|n|) \leq e^\tau} \frac{\chi(n)\bar{\psi}(n)}{n} = (1 - \chi(-1)\bar{\psi}(-1)) \sum_{P^+(n) \leq e^\tau} \frac{\chi(n)\bar{\psi}(n)}{n}.$$

Small unless  $\chi\bar{\psi}$  odd and  $\chi \approx \psi$ . So  $\psi$  induced by  $\xi$  and  $\chi\bar{\xi}$  odd. Then

$$\left| \sum_{P^+(|n|) \leq e^\tau} \frac{\chi(n)e(\frac{na}{b})}{n} \right| \sim \frac{2D^{\frac{1}{2}}}{b} \left| \sum_{D|d|b} \frac{\chi(\frac{b}{d})\mu(\frac{d}{D})\bar{\xi}(\frac{d}{D})}{\phi(d)/d} \sum_{P^+(n) \leq e^\tau}^{(n,d)=1} \frac{\chi(n)\bar{\xi}(n)}{n} \right|.$$

$$\implies 2e^{\gamma\tau} \lesssim \frac{2\sqrt{D}}{b} \sum_{D|d|b} \frac{d}{\phi(d)} \cdot \frac{\phi(d)}{d} e^{\gamma\tau} = \frac{2e^{\gamma\tau}(b/D)}{\sqrt{D}b/D} \leq \frac{2e^{\gamma\tau}}{\sqrt{D}}.$$

If  $D > 1$ , this is a contradiction. So  $\xi = 1$  and  $\chi$  is odd.

## Locating the maximum, continued

To summarize,

$$N(\chi) := \max_{\alpha \in [0,1]} \left| \sum_{P^+(|n|) \leq e^\tau} \frac{\chi(n)(1 - e(n\alpha))}{n} \right| > 2e^\gamma \tau \Rightarrow \chi \approx 1 \text{ and } \chi \text{ odd.}$$

Also, recall that  $\alpha_\chi$  location of max,  $|\alpha_\chi - a/b| < 1/(be^{\sqrt{\tau}})$ .

**Claim.**  $\exists c \leq \tau$  with  $|\sum_{P^+(n) \leq e^\tau, (n,c)=1} \chi(n)/n| \gtrsim e^\gamma \tau \phi(b)/b$ .

If  $b > \tau$ , we take  $c = 1$  (by Mont-Vaughan:  $N(\chi) \sim 2|\sum_{P^+(n) \leq e^\tau} \frac{\chi(n)}{n}|$ ).

If  $b \leq \tau$ , then we use a result of [Fouvry-Tenenbaum](#) on smooth numbers in APs to get asymptotics for the sum  $\sum_{n \leq N, P^+(n) \leq e^\tau} \chi(n)/n$ .

We then find that  $\exists N \in (e^{\sqrt{\tau}}, e^{\tau \log \tau}]$  (related to  $|\alpha_\chi - a/b|$ ) such that

$$\left| \sum_{\substack{n \leq N, (n,b)=1 \\ P^+(n) \leq e^\tau}} \frac{\chi(n)}{n} \right| + \left| \sum_{\substack{n > N, (n,b)=1 \\ P^+(n) \leq e^\tau}} \frac{\chi(n)}{n} \right| \gtrsim \frac{\phi(b)}{b} e^\gamma \tau.$$

## Lipschitz bounds for averages of $\chi$

$$S_1 = \sum_{\substack{n \leq N, (n,b)=1 \\ P^+(n) \leq e^\tau}} \frac{\chi(n)}{n}, \quad S_2 = \sum_{\substack{n > N, (n,b)=1 \\ P^+(n) \leq e^\tau}} \frac{\chi(n)}{n}.$$

We have  $|S_1| + |S_2| \gtrsim e^{\gamma\tau} \frac{\phi(b)}{b}$ , we want to show that  $|S_1 + S_2| \gtrsim e^{\gamma\tau} \frac{\phi(b)}{b}$ .

Note that  $|S_1| + |S_2| \lesssim e^{\gamma\tau} \phi(b)/b$ .

So, if  $S_j = \lambda_j |S_j|$  with  $|\lambda_j| = 1, j \in \{1, 2\}$ , then

$$0 \leq \sum_{\substack{n \leq N, (n,b)=1 \\ P^+(n) \leq e^\tau}} \frac{\chi(n) \overline{\lambda_1}}{n} + \sum_{\substack{n > N, (n,b)=1 \\ P^+(n) \leq e^\tau}} \frac{\chi(n) \overline{\lambda_2}}{n} = o(\tau \phi(b)/b).$$

So  $\chi(n) \sim \lambda_1$  for most  $n \leq N$  and  $\chi(n) \sim \lambda_2$  for most  $n > N$ .

Averages of mult. fncs vary slowly. Ideas from Halász's theorem +  $\chi \approx 1$ :

$$\frac{\sum_{n \leq x^{1+\delta}} \chi(n)}{x^{1+\delta}} - \frac{\sum_{n \leq x} \chi(n)}{x} \lesssim \delta \log(1/\delta) \quad (\delta \geq 1/\log x).$$

So  $\lambda_1 \sim \lambda_2$ , which implies that  $|S_1| + |S_2| \sim |S_1 + S_2|$ .

Thank you!