

Off-the-grid sparse estimation

Clarice Poon

University of Bath

Joint work with:

Nicolas Keriven, Gabriel Peyré, Mohammad Golbabaee

March 15, 2021

Introduction to the Blasso

Applying the Blasso to qMRI

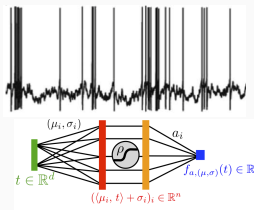
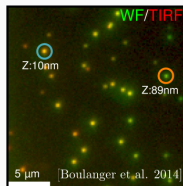
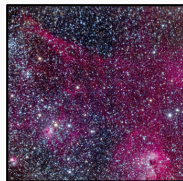
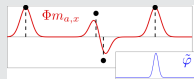
Problem description

Sparse linear models

Unknown sparse measure: $\mathbf{m}_{a,\theta} = \sum_{i=1}^n a_i \delta_{\theta_i}$ where $a_i \in \mathbb{R}$, $\theta_i \in \Theta \subset \mathbb{R}^d$.

Observe linear model: Define $\varphi : \Theta \rightarrow \mathcal{H}$ continuous

$$\Phi : \mathcal{M}(\Theta) \rightarrow \mathcal{H}, \quad \Phi \mathbf{m} \stackrel{\text{def.}}{=} \int_{\Theta} \varphi(\theta) d\mathbf{m}(\theta)$$



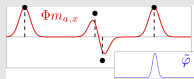
Problem description

Sparse linear models

Unknown sparse measure: $\mathbf{m}_{a,\theta} = \sum_{i=1}^n a_i \delta_{\theta_i}$ where $a_i \in \mathbb{R}$, $\theta_i \in \Theta \subset \mathbb{R}^d$.

Observe linear model: Define $\varphi : \Theta \rightarrow \mathcal{H}$ continuous

$$\Phi : \mathcal{M}(\Theta) \rightarrow \mathcal{H}, \quad \Phi \mathbf{m} \stackrel{\text{def.}}{=} \int_{\Theta} \varphi(\theta) d\mathbf{m}(\theta)$$



Fourier measurements: $\varphi(\theta) = (\exp(2\pi i l \theta))_{|l| \leq F} \in \mathbb{C}^{2F+1}$.

Then,

$$y = \Phi \mathbf{m} = \left(\sum_{i=1}^n a_i \exp(2\pi i l \theta_i) \right)_{|l| \leq F}.$$

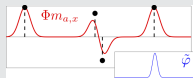
Problem description

Sparse linear models

Unknown sparse measure: $\mathbf{m}_{a,\theta} = \sum_{i=1}^n a_i \delta_{\theta_i}$ where $a_i \in \mathbb{R}$, $\theta_i \in \Theta \subset \mathbb{R}^d$.

Observe linear model: Define $\varphi : \Theta \rightarrow \mathcal{H}$ continuous

$$\Phi : \mathcal{M}(\Theta) \rightarrow \mathcal{H}, \quad \Phi \mathbf{m} \stackrel{\text{def.}}{=} \int_{\Theta} \varphi(\theta) d\mathbf{m}(\theta)$$



Deconvolution: $\varphi(\theta) = \kappa(\cdot - \theta) \in L^2(\mathbb{R}^d)$. Then,

$$y = \Phi \mathbf{m} = \sum_{i=1}^n a_i \kappa(\cdot - \theta_i).$$

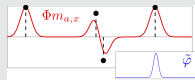
Problem description

Sparse linear models

Unknown sparse measure: $\mathbf{m}_{a,\theta} = \sum_{i=1}^n a_i \delta_{\theta_i}$ where $a_i \in \mathbb{R}$, $\theta_i \in \Theta \subset \mathbb{R}^d$.

Observe linear model: Define $\varphi : \Theta \rightarrow \mathcal{H}$ continuous

$$\Phi : \mathcal{M}(\Theta) \rightarrow \mathcal{H}, \quad \Phi \mathbf{m} \stackrel{\text{def.}}{=} \int_{\Theta} \varphi(\theta) d\mathbf{m}(\theta)$$



Laplace: $\varphi(\theta) = (\exp(-\theta t_k))_{k=1}^n \in \mathbb{R}^m$. Then,

$$y = \Phi \mathbf{m} = \left(\sum_{i=1}^n a_i \exp(-\theta_i t_k) \right)_{k=1}^m.$$

Multicompartment effects in imaging

We observe at voxel some time series measurement $y \in \mathbb{R}^T$

$$y = \sum_{i=1}^s a_i \varphi(\theta_i)$$

where $\varphi(\theta) \in \mathbb{R}^T$ models the behaviour of tissue type θ over time.

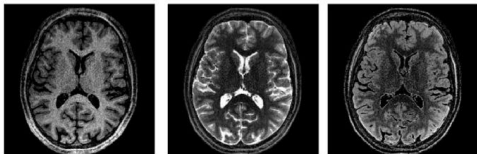


Figure 1: Contrast maps in quantitative MRI. Understanding multicompartment effects is important for accurate segmentation and studies of brain disorders.

The Beurling Lasso

Nonlinear least squares problem is **nonconvex**:

$$\min_{a, \theta} \frac{1}{2} \left\| \sum_i a_i \varphi(\theta_i) - y \right\|_2^2 + \lambda \|a\|_1$$

The Beurling Lasso

Nonlinear least squares problem is **nonconvex**:

$$\min_{a, \theta} \frac{1}{2} \left\| \sum_i a_i \varphi(\theta_i) - y \right\|_2^2 + \lambda \|a\|_1$$

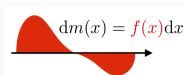
The Beurling Lasso

Minimisation over the space of measures is **convex**:

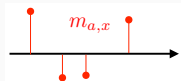
$$\min_{\mathbf{m} \in \mathcal{M}(\Theta)} \frac{1}{2} \|\Phi \mathbf{m} - y\|_2^2 + \lambda \|\mathbf{m}\|_{TV}. \quad (\mathcal{P}_\lambda(y))$$

[Beurling, '38, De Castro & Gamboa, '12, Bredies & Pikkarainen, '13]

$$\|\mathbf{m}\|_{TV} \stackrel{\text{def.}}{=} \sup_{\{\mathcal{A}_i\} \subset \Theta} \sum_i |\mathbf{m}(\mathcal{A}_i)|.$$



$$\|\mathbf{m}\|_{TV} = \|f\|_{L^1}$$



$$\|\mathbf{m}_{a,x}\|_{TV} = \|a\|_1$$

The Beurling Lasso

Nonlinear least squares problem is **nonconvex**:

$$\min_{a, \theta} \frac{1}{2} \left\| \sum_i a_i \varphi(\theta_i) - y \right\|_2^2 + \lambda \|a\|_1$$

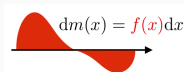
The Beurling Lasso

Minimisation over the space of measures is **convex**:

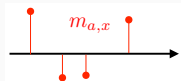
$$\min_{\mathbf{m} \in \mathcal{M}(\Theta)} \|\mathbf{m}\|_{TV} \quad \text{s.t.} \quad \Phi \mathbf{m} = y \quad (\mathcal{P}_0(y))$$

[Beurling, '38, De Castro & Gamboa, '12, Bredies & Pikkarainen, '13]

$$\|\mathbf{m}\|_{TV} \stackrel{\text{def.}}{=} \sup_{\{\mathcal{A}_i\} \subset \Theta} \sum_i |\mathbf{m}(\mathcal{A}_i)|.$$



$$\|\mathbf{m}\|_{TV} = \|f\|_{L^1}$$

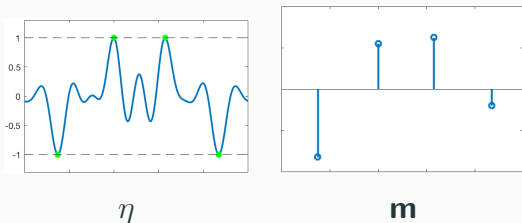


$$\|\mathbf{m}_{a,x}\|_{TV} = \|a\|_1$$

Key theoretical tool: dual certificates *

If you can find $\eta = \Phi^* p$ such that $|\eta(t)| < 1$ for all $t \notin \{\theta_i\}_i$ and $\eta(\theta_j) = \text{sign}(a_j)$, then

- Exact recovery of $\mathbf{m} = \sum_j a_j \delta_{\theta_j}$ from $y = \Phi \mathbf{m}$ by solving $\mathcal{P}_0(y)$.
- Stable recovery of $\mathbf{m} = \sum_j a_j \delta_{\theta_j}$ from $y = \Phi \mathbf{m} + w$ by solving $\mathcal{P}_\lambda(y)$, if in addition, $\text{sign}(a_j) \eta''(\theta_j) < 0$.



*Most Blasso papers make use of this result...

Key theoretical tool: dual certificates *

If you can find $\eta = \Phi^* p$ such that $|\eta(t)| < 1$ for all $t \notin \{\theta_i\}_i$ and $\eta(\theta_j) = \text{sign}(a_j)$, then

- Exact recovery of $\mathbf{m} = \sum_j a_j \delta_{\theta_j}$ from $y = \Phi \mathbf{m}$ by solving $\mathcal{P}_0(y)$.
- Stable recovery of $\mathbf{m} = \sum_j a_j \delta_{\theta_j}$ from $y = \Phi \mathbf{m} + w$ by solving $\mathcal{P}_\lambda(y)$, if in addition, $\text{sign}(a_j) \eta''(\theta_j) < 0$.

Minimal norm certificate

Most of the time, we look at

$$\eta_0 \stackrel{\text{def.}}{=} \operatorname{argmin}_{\eta = \Phi^* p} \|\mathbf{p}\| \quad \text{s.t.} \quad \begin{cases} \forall i, \eta(\theta_i) = \text{sign}(a_i) \\ \|\eta\|_\infty \leq 1. \end{cases}$$

*Most Blasso papers make use of this result...

Minimum separation

Candès & Fernandez-Granda, CPAM 2012

Consider $\varphi(\theta) = (e^{2\pi ik\theta})_{|k| \leq f_c}$. In dimension 1 and 2, η_0 is nondegenerate if

$$\Delta_\theta \stackrel{\text{def.}}{=} \min_{i \neq j} |\theta_i - \theta_j|_\infty \geq \frac{C}{f_c}$$

This result is **sharp**: If $\mathbf{m} = \delta_\theta - \delta_{\theta'}$ and $|\theta - \theta'| < \frac{1}{f_c}$, then no dual certificate exists (and this actually means that recovery is not possible).

Minimum separation

Candès & Fernandez-Granda, CPAM 2012

Consider $\varphi(\theta) = (e^{2\pi ik\theta})_{|k| \leq f_c}$. In dimension 1 and 2, η_0 is nondegenerate if

$$\Delta_\theta \stackrel{\text{def.}}{=} \min_{i \neq j} |\theta_i - \theta_j|_\infty \geq \frac{C}{f_c}$$

This result is **sharp**: If $\mathbf{m} = \delta_\theta - \delta_{\theta'}$ and $|\theta - \theta'| < \frac{1}{f_c}$, then no dual certificate exists (and this actually means that recovery is not possible).

*We first need to understand the minimum separation for arbitrary operators – need a **metric** to quantify what we mean by two spikes being close...*

Fisher-Rao distance

For $\theta, \theta' \in \Theta \subset \mathbb{R}^d$, define $K(\theta, \theta') \stackrel{\text{def.}}{=} \langle \varphi(\theta), \varphi(\theta') \rangle$.

Fisher metric:

$$\mathbf{g}_\theta = \nabla_1 \nabla_2 K(\theta, \theta) = [\nabla \varphi(\theta)][\nabla \varphi(\theta)]^\top \in \mathbb{R}^{d \times d}$$

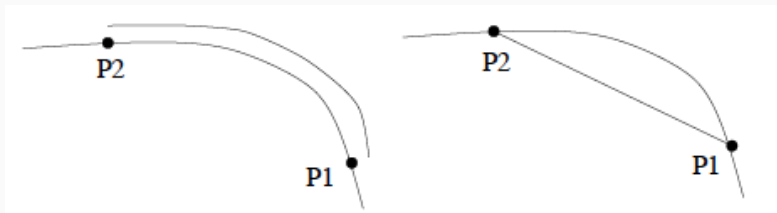
Fisher-Rao geodesic distance:

$$d_g(\theta, \theta') = \inf_{\gamma: \theta \rightarrow \theta'} \int_0^1 \sqrt{\langle \mathbf{g}_{\gamma(t)} \gamma'(t), \gamma'(t) \rangle} dt$$

Intuition

Statistical interpretation: If $\|\varphi(\theta)\| = 1$, then $(|\varphi(\theta)_i|^2)_i$ is a probability distribution.

Given $P_1 = \varphi(\theta)$ and $P_2 = \varphi(x')$:



The map $\theta \mapsto \varphi(\theta)$ embeds Θ into the sphere in \mathcal{H} and

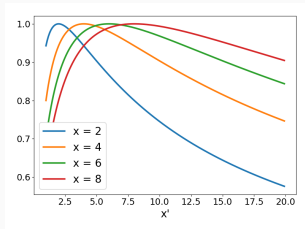
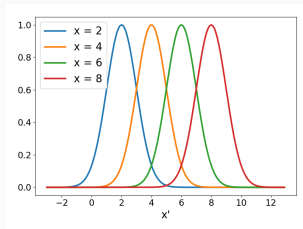
$$d_g(\theta, \theta') = \inf_{\gamma: \varphi(\theta) \rightarrow \varphi(\theta')} \int_0^1 \|\gamma'(t)\|_{\mathcal{H}} dt.$$

Examples

Fourier: $\Theta = \mathbb{T}^d$, $\varphi(\theta) = (e^{2\pi i k \theta})_{\|k\|_\infty \leq f_c}$ $\mathfrak{g}_\theta = f_c^2 \text{Id}$,
 $d_g(\theta, \theta') = f_c \|\theta - \theta'\|_2$.

Gaussian convolution: $\Theta = \mathbb{R}^d$
 $\varphi(\theta) \propto e^{-\|\theta - \cdot\|_\Sigma^2}$, $\mathfrak{g}_\theta = \Sigma$,
 $d_g(\theta, \theta') = \|\theta - \theta'\|_\Sigma$.

Laplace transform $\Theta = \mathbb{R}_+^d$,
 $\varphi(\theta) \propto e^{-\theta \cdot}$, $\mathfrak{g}_\theta = \text{diag}(1/\theta_i)$,
 $d_g(\theta, \theta') = \sqrt{\sum_i |\log(\theta_i/\theta'_i)|^2}$.



Recovery under minimum separation

Theorem (P., Keriven, Peyré '19)

Let $s \in \mathbb{N}$ and let $(\theta_i)_{i=1}^s$ be s.t. $\min_{i \neq j} d_g(\theta_i, \theta_j) \geq \Delta_{s,K}$.

Then: η_0 is nondegenerate.

Recovery under minimum separation

Theorem (P., Keriven, Peyré '19)

Let $s \in \mathbb{N}$ and let $(\theta_i)_{i=1}^s$ be s.t. $\min_{i \neq j} d_g(\theta_i, \theta_j) \geq \Delta_{s,K}$.

Then: η_0 is nondegenerate.

Examples:

Fourier coefficients: $\Delta = \min \left(\sqrt{d\sqrt{s}}, 2^d \right)$.

Gaussian deconvolution: $\Delta = \sqrt{\log(s)}$.

Laplace transform: $\Delta = d + \log(ds)$.

The separation distance $\Delta_{s,K}$ is independent of the problem parameters!

Compressed sensing result

E.g. $\varphi(\theta) = \left(e^{-2\pi i \omega_k \theta} \right)_{k=1}^m$ where $|k| \leq f_c$ are drawn randomly.

Compressed sensing result

E.g. $\varphi(\theta) = \left(e^{-2\pi i \omega_k \theta} \right)_{k=1}^m$ where $|k| \leq f_c$ are drawn randomly.

Setting:

Let (Ω, Λ) be a probability space and let $\varphi(\theta) = \left(\varphi_{\omega_k}(\theta) \right)_{k=1}^m$ where $\omega_k \stackrel{iid}{\sim} \Lambda$.

Consider recovery from $y = \Phi \left(\sum_{i=1}^s a_s \delta_{\theta_s} \right) + w$.

Compressed sensing result

E.g. $\varphi(\theta) = \left(e^{-2\pi i \omega_k \theta} \right)_{k=1}^m$ where $|k| \leq f_c$ are drawn randomly.

Setting:

Let (Ω, Λ) be a probability space and let $\varphi(\theta) = (\varphi_{\omega_k}(\theta))_{k=1}^m$ where $\omega_k \stackrel{iid}{\sim} \Lambda$.

Consider recovery from $y = \Phi \left(\sum_{i=1}^s a_s \delta_{\theta_s} \right) + w$.

Assumptions:

- Let $\theta \in \Theta^s$ be such that $\min_{j \neq k} d_g(\theta_j, \theta_k) \geq \Delta$.
- Let $\rho > 0$ and

$$m \geq C_{\bar{L}} \cdot s \cdot (\log^2(s/\rho) + \log(N^d/\rho))$$

where $C_{\bar{L}}$ and N depends on the derivatives of φ_{ω} and the domain diameter $\sup_{\theta, \theta' \in \Theta} d_g(\theta, \theta')$.

Compressed sensing result

Theorem (P., Keriven, Peyré '19)

Let $\lambda \sim \|w\| / \sqrt{s}$. With probability at least $1 - \rho$, any solution \mathbf{m} to $\mathcal{P}_\lambda(y)$ satisfies the following discrepancies to the true measures $\mathbf{m}_{a,\theta}$:

$$\max_{j=1}^s |a_j - \hat{a}_j| \lesssim s^{1/2} \|w\|. \quad \text{and} \quad \mathcal{T}_g^2(|\mathbf{m}|, |\mathbf{m}_{a,\theta}|) \lesssim s^{3/2} \|w\|.$$

where

$$\mathcal{T}_g^2(\mu, \nu) \stackrel{\text{def.}}{=} \inf_{\hat{\mu}, \hat{\nu}} W_g^2(\hat{\mu}, \hat{\nu}) + \|\mu - \hat{\mu}\|_{TV} + \|\nu - \hat{\nu}\|_{TV}.$$

Examples and remarks

Sampling Fourier coefficients with $\Theta = [0, 1]^d$:

$$m \sim d^2 \cdot s \cdot (\log^2(s) + \log(f_c^d))$$

Examples and remarks

Sampling Fourier coefficients with $\Theta = [0, 1]^d$:

$$m \sim d^2 \cdot s \cdot (\log^2(s) + \log(f_c^d))$$

Sampling the Laplace transform with $\Theta = (0, 1]^d$ with $\Lambda(t) \propto e^{-\alpha t}$ and $\alpha_i \sim d$:

$$m \sim d^6 \cdot s \cdot (\log^2(m) \log^2(s) + \log^4(m) \log(\log(m)^d))$$

Examples and remarks

Sampling Fourier coefficients with $\Theta = [0, 1]^d$:

$$m \sim d^2 \cdot s \cdot (\log^2(s) + \log(f_c^d))$$

Sampling the Laplace transform with $\Theta = (0, 1]^d$ with $\Lambda(t) \propto e^{-\alpha t}$ and $\alpha_j \sim d$:

$$m \sim d^6 \cdot s \cdot (\log^2(m) \log^2(s) + \log^4(m) \log(\log(m)^d))$$

Remark: Previous result by Tang et al. (2013) for sampling Fourier coefficients in 1D, **but** their result assumes that $\text{sign}(a_j)$ are distributed uniformly iid.

Introduction to the Blasso

Applying the Blasso to qMRI

Magnetic resonance imaging

MRI is one of the main applications of compressed sensing, this allows for subsampling

$$\min_{x \in \mathbb{R}^v} \lambda \|x\|_1 + \frac{1}{2} \|\mathcal{F}_\Omega x - y\|_2^2$$

Magnetic resonance imaging

MRI is one of the main applications of compressed sensing, this allows for subsampling

$$\min_{x \in \mathbb{R}^v} \lambda \|x\|_1 + \frac{1}{2} \|\mathcal{F}_\Omega x - y\|_2^2$$

Traditional MRI:

- The MR signal is obtained by applied the **same radio frequency (RF) pulse** repeatedly.
- $x \in \mathbb{R}^v$ is a **gray-valued image**, which captures the relative signal intensity changes between tissues, each image voxel is weighted by so-called T1,T2 values.

Quantitative MRI: measure T1, T2 values

Magnetic resonance fingerprinting [Ma et al '13, Nature] allowed this to be done in short clinically feasible scan times.

- Allow the **RF pulse** to vary over time.
- This results in a **time-series magnetisation images (TSMI)**

$$X = \begin{bmatrix} x_1 & x_2 & \cdots & x_v \end{bmatrix} \in \mathbb{R}^{T \times v}$$

with v voxels and T timeframes.

- The time dependent signal in each voxel is compared to a **dictionary of fingerprints** $\{\varphi(\theta_i)\}_i \subset \mathbb{R}^T$:
 - Precomputed by solving so-called **Bloch equations**,
 - Each fingerprint corresponds to $\theta = (T_1, T_2)$ values which depend on tissue type.

The quantitative MRI problem

Multicompartment effects: There can be more than one tissue type appearing in one image voxel.

TSMI with v voxels and T timeframes:

$$X = \begin{bmatrix} x_1 & x_2 & \cdots & x_v \end{bmatrix} \in \mathbb{R}^{T \times v}$$

For each $x = x_i$,

$$x = \sum_s c_s \varphi(\theta_s) \in \mathbb{R}^T$$

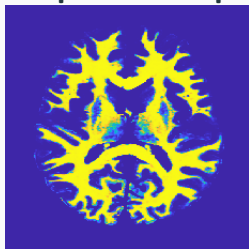
- $c_s \geq 0$ are mixture weights
- $\varphi : \Theta \rightarrow \mathbb{R}^T$ is the Bloch magnetisation response model.
- Θ is the domain of NMR properties.

Visualisation

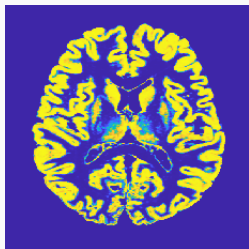
TSMI:



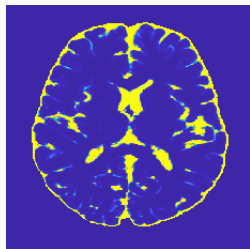
Component maps:



$$\theta_1 = (784, 77)$$



$$\theta_2 = (1216, 96)$$



$$\theta_3 = (4083, 1394)$$

The quantitative MRI problem

Previous approaches:

- discretize the domain Θ , to $\{\theta_s\}_{s=1}^N$, form a dictionary

$$D_\theta \stackrel{\text{def.}}{=} \begin{bmatrix} \varphi(\theta_1) & \varphi(\theta_2) & \cdots & \varphi(\theta_N) \end{bmatrix}$$

- solve for $C \in \mathbb{R}^{\nu \times N}$,

$$D_\theta C^\top = X$$

where each column $C_s \in \mathbb{R}^\nu$ correspond to the θ_s dependent mixture weights across all voxels.

Solve problem of form:

$$\min_C \frac{1}{2} \|D_\theta C^\top - X\|_F^2 + J(C)$$

Formulation as Blasso with vector-valued measures

Write $\mathbf{m} = \sum_{s=1}^k \mathbf{C}_s^\top \delta_{\theta_s} \in \mathcal{M}(\Theta; \mathbb{R}^v)$.

Then, we have

$$\Phi \mathbf{m} = \int \varphi(\theta) d\mathbf{m}(\theta) = \sum_s \varphi(\theta_s) \mathbf{C}_s^\top.$$

Formulation as Blasso with vector-valued measures

Write $\mathbf{m} = \sum_{s=1}^k C_s^\top \delta_{\theta_s} \in \mathcal{M}(\Theta; \mathbb{R}^v)$.

Then, we have

$$\Phi \mathbf{m} = \int \varphi(\theta) d\mathbf{m}(\theta) = \sum_s \varphi(\theta_s) C_s^\top.$$

Remark: Other models are possible, e.g.

$$\mathbf{m} = \sum_{s=1}^k C_s^\top g_\sigma(\theta_s - \cdot).$$

Then,

$$\Phi \mathbf{m} = \langle \varphi \star g_\sigma, \sum_s C_s^\top \delta_{\theta_s} \rangle$$

Formulation as Blasso with vector-valued measures

Write $\mathbf{m} = \sum_{s=1}^k \mathbf{C}_s^\top \delta_{\theta_s} \in \mathcal{M}(\Theta; \mathbb{R}^v)$.

Then, we have

$$\Phi \mathbf{m} = \int \varphi(\theta) d\mathbf{m}(\theta) = \sum_s \varphi(\theta_s) \mathbf{C}_s^\top.$$

Total variation of vector valued measures

If a measure takes values in a normed space \mathcal{V} endowed with norm $\|\cdot\|_{\mathcal{V}}$, then define

$$|\mathbf{m}|_{\mathcal{V}} = \sup_{\{\mathcal{A}_j\} \subset \mathcal{V}} \sum_{j=1}^N \|\mathbf{m}(\mathcal{A}_j)\|_{\mathcal{V}}.$$

We need to choose $\|\cdot\|_{\mathcal{V}}$.

Sparse-group-Blasso

We consider regularisation with the following mixed norm:

$$\|\mathbf{m}\|_{\beta} \stackrel{\text{def.}}{=} (1 - \beta) \|\mathbf{m}\|_1 + \beta \sqrt{v} \|\mathbf{m}\|_2.$$

So:

$$\min_{\mathbf{m} \in \mathcal{M}(\Theta; \mathbb{R}^v)} \lambda \|\mathbf{m}\|_{\beta} + \frac{1}{2} \|\mathbf{X} - \Phi \mathbf{m}\|_F^2.$$

NB: $\|\sum_s C_s \delta_s\|_{\beta} = (1 - \beta) \sum_s \|C_s\|_1 + \beta \sqrt{v} \sum_s \|C_s\|_2.$

- $\sum_s \|C_s\|_2$ enforces group sparsity.
- $\sum_s \|C_s\|_1$ enforces sparsity within each mixture map.

This is the continuous counterpart of the *sparse-group lasso* [Simon, Hastie & Tibshirani, JCGS, 2013].

Conditional gradient descent

Solve $\min_{x \in \mathcal{C}} f(x)$, \mathcal{C} is a compact convex set in Banach space:

$$y_t \in \operatorname{argmin}_{y \in \mathcal{C}} \nabla f(x_t)^\top y$$

$$x_{t+1} = (1 - \gamma_t)x_t + \gamma_t y_t$$

Conditional gradient descent

Solve $\min_{x \in \mathcal{C}} f(x)$, \mathcal{C} is a compact convex set in Banach space:

$$y_t \in \operatorname{argmin}_{y \in \mathcal{C}} \nabla f(x_t)^\top y$$

$$x_{t+1} = (1 - \gamma_t)x_t + \gamma_t y_t$$

For our problem $\min_{\mathbf{m}} \lambda \|\mathbf{m}\|_{TV} + \frac{1}{2} \|\Phi \mathbf{m} - X\|_F^2$:

$$\lambda \|\mathbf{m}\|_\beta \leq \|0\|_{TV} + \frac{1}{2} \|\Phi 0 - X\|_F^2 = \|X\|_F^2 / 2.$$

Therefore, we solve

$$\min_{t, \mathbf{m} \in \mathcal{C}} f(t) = \lambda t + \frac{1}{2} \|X - \Phi \mathbf{m}\|_F^2$$

where $\mathcal{K} = \left\{ (t, \mathbf{m}) \in \mathbb{R}_+ \times \mathcal{M} ; \|\mathbf{m}\|_\beta \leq t \leq \|X\|_F^2 / (2\lambda) \right\}$.

Convergence of objective is $\mathcal{O}(1/k)$ with k being iteration.

Frank-Wolfe/Conditional gradient iterations[†]

Inputs: TSMI X , Bloch model $\varphi(\cdot)$, params $\alpha, \beta > 0$.

Outputs: NMR parameters θ , mixture weights C .

Initialise: $i = 0$, $\theta^0 = \{\}$, $C^0 = \{\}$, $\eta^0 = \frac{1}{\alpha} \Phi^* X$.

repeat

[†]Follows [Denoyelle et al, Inverse problems '19]

Frank-Wolfe/Conditional gradient iterations[†]

Inputs: TSMI X , Bloch model $\varphi(\cdot)$, params $\alpha, \beta > 0$.

Outputs: NMR parameters θ , mixture weights C .

Initialise: $i = 0$, $\theta^0 = \{\}$, $C^0 = \{\}$, $\eta^0 = \frac{1}{\alpha} \Phi^* X$.

repeat

Let $\theta \in \operatorname{argmax}_{\theta \in \Theta} \sum_{s=1}^v (\eta^i(\theta)_s - (1 - \beta))_+^2$

[†]Follows [Denoyelle et al, Inverse problems '19]

Frank-Wolfe/Conditional gradient iterations[†]

Inputs: TSMI X , Bloch model $\varphi(\cdot)$, params $\alpha, \beta > 0$.

Outputs: NMR parameters θ , mixture weights C .

Initialise: $i = 0$, $\theta^0 = \{\}$, $C^0 = \{\}$, $\eta^0 = \frac{1}{\alpha} \Phi^* X$.

repeat

Let $\theta \in \operatorname{argmax}_{\theta \in \Theta} \sum_{s=1}^v (\eta^i(\theta)_s - (1 - \beta))_+^2$

$\theta^{i+\frac{1}{2}} = \theta^i \cup \{\theta\}$

[†]Follows [Denoyelle et al, Inverse problems '19]

Frank-Wolfe/Conditional gradient iterations[†]

Inputs: TSMI X , Bloch model $\varphi(\cdot)$, params $\alpha, \beta > 0$.

Outputs: NMR parameters θ , mixture weights C .

Initialise: $i = 0$, $\theta^0 = \{\}$, $C^0 = \{\}$, $\eta^0 = \frac{1}{\alpha} \Phi^* X$.

repeat

Let $\theta \in \operatorname{argmax}_{\theta \in \Theta} \sum_{s=1}^v (\eta^i(\theta)_s - (1 - \beta))_+^2$

$\theta^{i+\frac{1}{2}} = \theta^i \cup \{\theta\}$

$C^{i+\frac{1}{2}} \in \operatorname{argmin}_{C \in \mathbb{R}_+^{k \times v}} \frac{1}{2} \left\| X - D_{\theta^{i+\frac{1}{2}}} C \right\|_F^2 + \alpha \|C\|_\beta$

[†]Follows [Denoyelle et al, Inverse problems '19]

Frank-Wolfe/Conditional gradient iterations[†]

Inputs: TSMI X , Bloch model $\varphi(\cdot)$, params $\alpha, \beta > 0$.

Outputs: NMR parameters θ , mixture weights C .

Initialise: $i = 0$, $\theta^0 = \{\}$, $C^0 = \{\}$, $\eta^0 = \frac{1}{\alpha} \Phi^* X$.

repeat

Let $\theta \in \operatorname{argmax}_{\theta \in \Theta} \sum_{s=1}^v (\eta^i(\theta)_s - (1 - \beta))_+^2$

$\theta^{i+\frac{1}{2}} = \theta^i \cup \{\theta\}$

$C^{i+\frac{1}{2}} \in \operatorname{argmin}_{C \in \mathbb{R}_+^{k \times v}} \frac{1}{2} \left\| X - D_{\theta^{i+\frac{1}{2}}} C \right\|_F^2 + \alpha \|C\|_\beta$

Initialising with $C^{i+1/2}$ and $\theta^{i+1/2}$, solve

$$(C^{i+1}, \theta^{i+1}) \in \operatorname{argmin}_{\theta, C} \frac{1}{2} \left\| X - D_\theta C^\top \right\|_F^2 + \alpha \|C\|_\beta$$

[†]Follows [Denoyelle et al, Inverse problems '19]

Frank-Wolfe/Conditional gradient iterations[†]

Inputs: TSMI X , Bloch model $\varphi(\cdot)$, params $\alpha, \beta > 0$.

Outputs: NMR parameters θ , mixture weights C .

Initialise: $i = 0$, $\theta^0 = \{\}$, $C^0 = \{\}$, $\eta^0 = \frac{1}{\alpha} \Phi^* X$.

repeat

Let $\theta \in \operatorname{argmax}_{\theta \in \Theta} \sum_{s=1}^v (\eta^i(\theta)_s - (1 - \beta))_+^2$

$\theta^{i+\frac{1}{2}} = \theta^i \cup \{\theta\}$

$C^{i+\frac{1}{2}} \in \operatorname{argmin}_{C \in \mathbb{R}_+^{k \times v}} \frac{1}{2} \left\| X - D_{\theta^{i+\frac{1}{2}}} C \right\|_F^2 + \alpha \|C\|_\beta$

Initialising with $C^{i+1/2}$ and $\theta^{i+1/2}$, solve

$$(C^{i+1}, \theta^{i+1}) \in \operatorname{argmin}_{\theta, C} \frac{1}{2} \left\| X - D_\theta C^\top \right\|_F^2 + \alpha \|C\|_\beta$$

Define $\eta^{i+1} = \frac{1}{\alpha} \Phi^*(X - D_{\theta^{i+1}}(C^{i+1})^\top)$

[†]Follows [Denoyelle et al, Inverse problems '19]

Frank-Wolfe/Conditional gradient iterations[†]

Inputs: TSMI X , Bloch model $\varphi(\cdot)$, params $\alpha, \beta > 0$.

Outputs: NMR parameters θ , mixture weights C .

Initialise: $i = 0$, $\theta^0 = \{\}$, $C^0 = \{\}$, $\eta^0 = \frac{1}{\alpha} \Phi^* X$.

repeat

Let $\theta \in \operatorname{argmax}_{\theta \in \Theta} \sum_{s=1}^v (\eta^i(\theta)_s - (1 - \beta))_+^2$

$\theta^{i+\frac{1}{2}} = \theta^i \cup \{\theta\}$

$C^{i+\frac{1}{2}} \in \operatorname{argmin}_{C \in \mathbb{R}_+^{k \times v}} \frac{1}{2} \left\| X - D_{\theta^{i+\frac{1}{2}}} C \right\|_F^2 + \alpha \|C\|_\beta$

Initialising with $C^{i+1/2}$ and $\theta^{i+1/2}$, solve

$$(C^{i+1}, \theta^{i+1}) \in \operatorname{argmin}_{\theta, C} \frac{1}{2} \left\| X - D_\theta C^\top \right\|_F^2 + \alpha \|C\|_\beta$$

Define $\eta^{i+1} = \frac{1}{\alpha} \Phi^*(X - D_{\theta^{i+1}}(C^{i+1})^\top)$

$i = i + 1$

until $\sup_{\theta \in \mathcal{T}} \sum_{s=1}^v (\eta_s^i(\theta) - (1 - \beta))_+^2 \leq v\beta^2$

[†]Follows [Denoyelle et al, Inverse problems '19]

Setup for numerics

The MRF data came from a healthy volunteer's brain, a variable density spiral trajectory was used for k -space sampling.

- MRF excitation sequences with $T = 1000$ timepoints. That is $\varphi(\theta) \in \mathbb{R}^T$. Acquisition window around 10s.
- The number of image voxels per timeframe is 230x230.
- First recover the TSMI from k -space measurements using LRTV. This is standard compressed sensing with TV regularisation.
- We then apply SGB-Lasso to recover mixture maps.

Typically, TSMI is complex valued, however, it is often assume to have constant-valued phase which can be subtracted and removed.

Useful because positivity constraint helps in practice.

‡Jiang et al., MRI, 2015; Nagtegaal et al., Magnetic resonance in medicine, 2020.

It is observed that Block responses have low rank approximations

$$\varphi(\theta) \approx VV^T\varphi(\theta)$$

where $V \in \mathbb{R}^{T \times \tau}$ with $\tau \ll T$ (we took $\tau = 10$) and the columns of V form an orthonormal system.

This V comes from PCA of a large simulated dictionary.

So, instead, work with $\tilde{\varphi} = V^T\varphi(\theta) \in \mathbb{R}^\tau$ and $\tilde{X} = V^T X$.

§McGivney et al. IEEE TMI (2014). Cline et al. MRI (2017)

Instead of working with $\tilde{\varphi} = V^T \varphi(\theta) \in \mathbb{R}^\tau$, train a 2 layer neural network

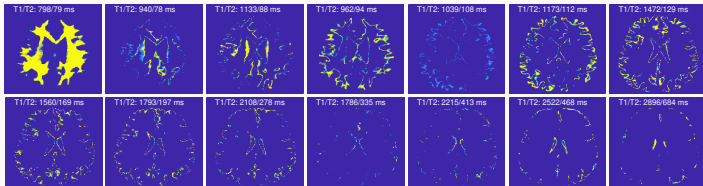
$$\mathcal{N} : \theta \in \Theta \mapsto \tilde{\varphi}(\theta).$$

This means that $\tilde{\varphi}$ and its Jacobian can be evaluated efficiently.

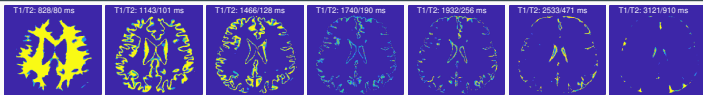
¶Chen et al, MICCAI (2020); Gómez et al, Scientific reports (2020)

Effects of β

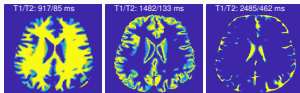
$\beta = 0.0001$



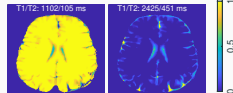
$\beta = 0.001$



$\beta = 0.01$



$\beta = 0.05$



Comparison against baseline methods:

- PVMRF ^{||}. Estimate dictionary using k -means
- SPIJN ^{**} Group sparsity regularization.
- BayesianMRF ^{††} Enforces sparsity.

		T1 (ms)			
Tissue	Literature	SGBlasso	PVMRF	SPIJN	BayesianMRF
WM	694 – 862	829	806	699	821
GM	1074 – 1174	1114	1165	1483	874

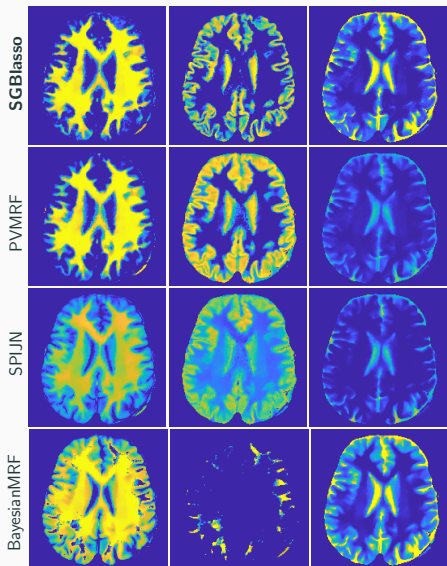
		T2 (ms)			
Tissue	Literature	SGBlasso	PVMRF	SPIJN	BayesianMRF
WM	68 – 87	81	80	51	77
GM	87 – 103	102	105	164	82

^{||}Deshmane et al, NMR in Biomedicine, 2018

^{**}Nagtegaal et al, Magnetic resonance in medicine 2020

^{††}McGivney et al, Magnetic resonance in medicine 2018

Comparison with existing methods



Summary

- Introduction of the Fisher metric, which offers a way of imposing the separation condition. This provides a unified way of approaching nontranslational invariant problems.
- The Blasso framework gives promising results for the problem of multi-compartment analysis in MRF.

Papers:

- *The geometry of off-the-grid compressed sensing*, P., Keriven & Peyré, [arXiv:1802.08464](https://arxiv.org/abs/1802.08464)
- *An off-the-grid approach to multi-compartment magnetic resonance fingerprinting*, Golbabaee & P., [arXiv:2011.11193](https://arxiv.org/abs/2011.11193)

Summary

- Introduction of the Fisher metric, which offers a way of imposing the separation condition. This provides a unified way of approaching nontranslational invariant problems.
- The Blasso framework gives promising results for the problem of multi-compartment analysis in MRF.

Papers:

- *The geometry of off-the-grid compressed sensing*, P., Keriven & Peyré, [arXiv:1802.08464](https://arxiv.org/abs/1802.08464)
- *An off-the-grid approach to multi-compartment magnetic resonance fingerprinting*, Golbabaee & P., [arXiv:2011.11193](https://arxiv.org/abs/2011.11193)

Thanks for listening